# Lecture 8 The Extended Kalman filter

- Nonlinear filtering
- Extended Kalman filter
- Linearization and random variables

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### Nonlinear filtering

• nonlinear Markov model:

 $x(t+1) = f(x(t), w(t)), \qquad y(t) = g(x(t), v(t))$ 

- f is (possibly nonlinear) dynamics function - g is (possibly nonlinear) measurement or output function -  $w(0), w(1), \ldots, v(0), v(1), \ldots$  are independent - even if w, v Gaussian, x and y need not be

• nonlinear filtering problem: find, e.g.,

$$\hat{x}(t|t-1) = \mathbf{E}(x(t)|y(0), \dots, y(t-1)), \qquad \hat{x}(t|t) = \mathbf{E}(x(t)|y(0), \dots, y(t))$$

• general nonlinear filtering solution involves a PDE, and is not practical

- extended Kalman filter (EKF) is *heuristic* for nonlinear filtering problem
- often works well (when tuned properly), but sometimes not
- widely used in practice
- based on
  - linearizing dynamics and output functions at current estimate
    propagating an approximation of the conditional expectation and covariance

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## Linearization and random variables

- consider  $\phi : \mathbf{R}^n \to \mathbf{R}^m$
- suppose  $\mathbf{E} x = \bar{x}$ ,  $\mathbf{E}(x \bar{x})(x \bar{x})^T = \Sigma_x$ , and  $y = \phi(x)$
- if  $\Sigma_x$  is small,  $\phi$  is not too nonlinear,

$$y \approx \tilde{y} = \phi(\bar{x}) + D\phi(\bar{x})(x - \bar{x})$$
$$\tilde{y} \sim \mathcal{N}(\phi(\bar{x}), D\phi(\bar{x})\Sigma_x D\phi(\bar{x})^T)$$

• gives *approximation* for mean and covariance of nonlinear function of random variable:

$$\bar{y} \approx \phi(\bar{x}), \qquad \Sigma_y \approx D\phi(\bar{x})\Sigma_x D\phi(\bar{x})^T$$

• if  $\Sigma_x$  is not small compared to 'curvature' of  $\phi$ , these estimates are poor

• a good estimate can be found by Monte Carlo simulation:

$$\bar{y} \approx \bar{y}^{\mathrm{mc}} = \frac{1}{N} \sum_{i=1}^{N} \phi(x^{(i)})$$
  
$$\Sigma_{y} \approx \frac{1}{N} \sum_{i=1}^{N} \left( \phi(x^{(i)}) - \bar{y}^{\mathrm{mc}} \right) \left( \phi(x^{(i)}) - \bar{y}^{\mathrm{mc}} \right)^{T}$$

where  $x^{(1)}, \ldots, x^{(N)}$  are samples from the distribution of x, and N is large

• another method: use Monte Carlo formulas, with a small number of nonrandom samples chosen as 'typical', e.g., the 90% confidence ellipsoid semi-axis endpoints

$$x^{(i)} = \bar{x} \pm \beta v_i, \qquad \Sigma_x = V \Lambda V^T$$

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#### Example

 $x \sim \mathcal{N}(0, 1), y = \exp(x)$ 

(for this case we can compute mean and variance of y exactly)

|   | $ar{y}$           | $\sigma_y$               |
|---|-------------------|--------------------------|
| exact values  | $e^{1/2} = 1.649$ | $\sqrt{e^2 - e} = 2.161$ |
| linearization   | 1.000             | 1.000                    |
| Monte Carlo ( $N = 10$ )                                | 1.385             | 1.068                    |
| Monte Carlo ( $N = 100$ )                               | 1.430             | 1.776                    |
| Sigma points $(x = \bar{x}, \ \bar{x} \pm 1.5\sigma_x)$ | 1.902             | 2.268                    |

#### **Extended Kalman filter**

- initialization:  $\hat{x}(0|-1) = \bar{x}_0$ ,  $\Sigma(0|-1) = \Sigma_0$
- measurement update
  - linearize output function at  $x = \hat{x}(t|t-1)$ :

$$C = \frac{\partial g}{\partial x}(\hat{x}(t|t-1), 0)$$
$$V = \frac{\partial g}{\partial v}(\hat{x}(t|t-1), 0)\Sigma_v \frac{\partial g}{\partial v}(\hat{x}(t|t-1), 0)^T$$

- measurement update based on linearization

$$\hat{x}(t|t) = \hat{x}(t|t-1) + \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} \dots \\ \dots \left( y(t) - g(\hat{x}(t|t-1), 0) \right)$$
$$\Sigma_{t|t} = \sum_{t|t-1} \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C \sum_{t|t-1} C^T \left( C \sum_{t|t-1} C^T + V \right)^{-1} C \sum_{t|t-1} C$$

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- time update
  - linearize dynamics function at  $x = \hat{x}(t|t)$ :

$$A = \frac{\partial f}{\partial x}(\hat{x}(t|t), 0)$$
$$W = \frac{\partial f}{\partial w}(\hat{x}(t|t), 0)\Sigma_w \frac{\partial f}{\partial w}(\hat{x}(t|t), 0)^T$$

- time update based on linearization

$$\hat{x}(t+1|t) = f(\hat{x}(t|t), 0), \qquad \Sigma_{t+1|t} = A\Sigma_{t|t}A^T + W$$

- replacing linearization with Monte Carlo yields particle filter
- replacing linearization with sigma-point estimates yields *unscented Kalman filter* (UKF)

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#### Example

- $p(t),\,u(t)\in {\bf R}^2$  are position and velocity of vehicle, with  $(p(0),u(0))\sim \mathcal{N}(0,I)$
- vehicle dynamics:

$$p(t+1) = p(t) + 0.1u(t),$$
  $u(t+1) = \begin{bmatrix} 0.85 & 0.15\\ -0.1 & 0.85 \end{bmatrix} u(t) + w(t)$ 

w(t) are IID  $\mathcal{N}(0, I)$ 

• measurements: noisy measurements of distance to 9 points  $p_i \in \mathbf{R}^2$ 

$$y_i(t) = ||p(t) - p_i|| + v_i(t), \quad i = 1, \dots, 9,$$

 $v_i(t)$  are IID  $\mathcal{N}(0, 0.3^2)$ 

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### **EKF** results

- EKF initialized with  $\hat{x}(0|-1)=0,$   $\Sigma(0|-1)=I,$  where x=(p,u)
- $p_i$  shown as stars; p(t) as dotted curve;  $\hat{p}(t|t)$  as solid curve



## Current position estimation error



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Current position estimation predicted error



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